

THE LOG-NORMAL MODEL OF RELATIONAL DATA

Let Z_{jk} be the j^{th} stimulus coordinate on the k^{th} dimension, $k=1, \dots, s$, where s is the number of dimensions. Let d_{jm} be the Euclidean distance between stimulus j and stimulus m in the s -dimensional space:

$$d_{jm} = \sqrt{\sum_{k=1}^s (Z_{jk} - Z_{mk})^2} \quad (1)$$

We assume that the observed distances, d_{jm}^* , are

$$d_{jm}^* = d_{jm} + \varepsilon_{jm} \quad (2)$$

And that they are drawn from the log-normal distribution because distances are inherently positive:

$$\ln(d_{jm}^*) \sim N(\ln(d_{jm}), \sigma^2) \quad (3)$$

That is

$$f(d_{jm}^*) = \frac{1}{(2\pi\sigma^2)^{\frac{1}{2}} d_{jm}^*} e^{\left(-\frac{1}{2\sigma^2}(\ln(d_{jm}^*) - \ln(d_{jm}))^2\right)} \quad (4)$$

The log-normal is a more realistic model of the noise process because, by definition of the log-normal, $d_{jm}^* > 0$ so that $(d_{jm} + \varepsilon_{jm}) > 0$ and $\sigma > 0$, where σ is a *shape parameter*. The mean and variance are:

$$E[d_{jm}^*] = d_{jm} e^{\frac{1}{2}\sigma^2} \quad \text{and} \quad \text{VAR}(d_{jm}^*) = (e^{\sigma^2} - 1) d_{jm}^2 e^{\sigma^2} \quad (5)$$

So that as $d_{jm} \rightarrow 0$ $E(d_{jm}^*) \rightarrow 0$ and $\text{VAR}(d_{jm}^*) \rightarrow 0$. The upshot is that the smaller the observed distance the smaller the variance of that distance because $d_{jm}^* = d_{jm} + \varepsilon_{jm}$.

Our likelihood function is:

$$L^*(Z_{jk} | \mathcal{D}^*) = \frac{1}{(2\pi\sigma^2)^{\frac{q(q-1)/2}{2}}} \left(\prod_{j=1}^{q-1} \prod_{m=j+1}^q \frac{1}{d_{jm}^*} \right) e^{-\frac{1}{2\sigma^2} \sum_{j=1}^{q-1} \sum_{m=j+1}^q \left(\ln(d_{jm}^*) - \ln \left(\sqrt{\sum_{k=1}^s (Z_{jk} - Z_{mk})^2} \right) \right)^2} \quad (6)$$

To implement a Bayesian model we use simple normal prior distributions for the stimuli coordinates:

$$\xi(Z_{jk}) = \frac{1}{(2\pi\kappa^2)^{\frac{1}{2}}} e^{-\frac{Z_{jk}^2}{2\kappa^2}} \quad (7)$$

and a uniform prior for the variance term:

$$\xi(\sigma^2) = \frac{1}{c}, \quad 0 < c < b \quad (8)$$

Where, empirically, b is no greater than 2.

Hence, our posterior distribution is:

$$\xi(Z_{jk} | \mathcal{D}^*) \propto \prod_{j=1}^{q-1} \prod_{m=j+1}^q \left\{ f_{jm}(Z_{jm} | d_{jm}^*) \right\} \xi(Z_{11}) \xi(Z_{12}) \dots \xi(Z_{1s}) \xi(Z_{21}) \dots \xi(Z_{qs}) \xi(\sigma^2) \quad (9)$$

Taking the log of the right hand side and dropping the unnecessary constants:

$$\begin{aligned}
\ell n \xi \propto & -\frac{q(q-1)/2}{2} \ln(\sigma^2) - \frac{1}{2\sigma^2} \sum_{j=1}^{q-1} \sum_{m=j+1}^q \left(\ln(d_{jm}^*) - \ln \left(\sqrt{\sum_{k=1}^s (Z_{jk} - Z_{mk})^2} \right) \right)^2 \\
& - \frac{1}{2\kappa^2} \left(\sum_{j=1}^q \sum_{k=1}^s Z_{jk}^2 \right) - \ln(c) = \\
& -\frac{q(q-1)/2}{2} \ln(\sigma^2) - \frac{1}{2\sigma^2} \sum_{j=1}^{q-1} \sum_{m=j+1}^q \left(\ln(d_{jm}^*) - \ln(d_{jm}) \right)^2 - \frac{1}{2\kappa^2} \left(\sum_{j=1}^q \sum_{k=1}^s Z_{jk}^2 \right) - \ln(c) \quad (10)
\end{aligned}$$